# Obesity, inflammatory markers, and endometrial cancer risk: a prospective case-control study 

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#### Abstract

Obesity, a major risk factor for endometrial cancer, is a low-grade inflammatory state characterized by elevated concentrations of cytokines and acute phase reactants. The current study had two aims: first to investigate the associations of C-reactive protein (CRP), interleukin 6 (IL6), and IL1 receptor antagonist (IL1Ra) with endometrial cancer risk and second to examine to which extent these markers can influence the association between obesity and endometrial cancer. We conducted a case-control study, nested within the European Prospective Investigation into Cancer and Nutrition, which comprised 305 incident cases of endometrial cancer and 574 matched controls. CRP, IL6, and IL1Ra were measured in prospectively collected blood specimens by immunoassays. Data were analyzed using conditional logistic regression. All statistical tests were two-sided, and $P$ values $<0.05$ were considered statistically significant. We observed a significant increase in risk of endometrial cancer with elevated levels of CRP (odds ratio (OR) for top versus bottom quartile: $1.58,95 \%$ confidence interval (CI): 1.03-2.41, $P_{\text {trend }}=0.02$ ), IL6 (OR for top versus bottom quartile: $1.66,95 \% \mathrm{CI}: 1.08-2.54, P_{\text {trend }}=0.008$ ), and IL1Ra (OR for top versus bottom quartile: $1.82,95 \% \mathrm{CI}: 1.22-2.73, P_{\text {trend }}=0.004$ ). After adjustment for body mass index (BMI), the estimates were strongly reduced and became non-significant. The association between BMI and endometrial cancer was also substantially attenuated ( $\sim 10-20 \%$ ) after adjustment for inflammatory markers, even when the effects of C-peptide or estrone had already been taken into account. We provided epidemiological evidence that chronic inflammation might mediate the association between obesity and endometrial cancer and that endometrial carcinogenesis could be promoted by an inflammatory milieu.


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## Introduction

Excess body weight is one of the strongest risk factors for endometrial cancer and accounts for about $50 \%$ of the cases in Europe and in the USA (International Agency for Research on Cancer (IARC) 2002, Calle \& Kaaks 2004, World Cancer Research Fund/American Institute for Cancer Research 2007). Obese women (body mass index $(\mathrm{BMI})>30 \mathrm{~kg} / \mathrm{m}^{2}$ ) have a threefold increased risk of developing endometrial cancer compared to normal weight women (Calle \& Kaaks 2004, Reeves et al. 2007).

Several biological mechanisms could mediate the association between obesity and endometrial cancer risk. In postmenopausal women, excess weight leads, through peripheral conversion of androgens, to increased circulating bioavailable estrogens which in turn, when not counterbalanced by progesterone, increases the mitogenic activity of endometrial cells (Key \& Pike 1988). In premenopausal women, obesity has been associated with the development of polycystic ovary syndrome, a disease that affects $4-8 \%$ of women and is characterized by ovarian hyperandrogenism, anovulation, and progesterone deficiency (Ehrmann et al. 1995, Franks 1995). Obesity-associated
hyperinsulinemia might also play an important role in endometrial carcinogenesis, in both pre and postmenopausal women, either directly, by stimulating endometrial cell proliferation, or indirectly, through the sex steroid and the insulin-like growth factor 1 pathways (Kaaks et al. 2002).

More recently, it has been proposed that other adiposity-related factors such as cytokines and adipokines might contribute to endometrial cancer initiation and progression and to the obesity-related increase in risk (Modugno et al. 2005). Excess body weight is associated with a systemic low-grade inflammatory condition (Das 2001), characterized by elevations in circulating pro-inflammatory cytokines and acute phase proteins (Mohamed-Ali et al. 1997, Visser et al. 1999, Yudkin et al. 1999, Meier et al. 2002). Chronic inflammation promotes angiogenesis, sustains cell proliferation, and increases production of free radicals that cause DNA damage, leading to tumor initiation and development (Coussens \& Werb 2002). Inflammatory processes also play a central role in the regulation of endometrial mucosa growth and shedding during the menstrual cycle (Kelly et al. 2001) as well as in endometrial repair following menstruation (Salamonsen 2003).

C-reactive protein (CRP) is an acute phase protein produced by the liver in response to tissue damage and inflammation. CRP binds to damaged cellular tissues and components to facilitate phagocytosis and clearance by macrophages and leukocytes (Pasche \& Serhan 2004). The major stimulus of CRP synthesis is interleukin 6 (IL6; Goldman \& Liu 1987, Li \& Goldman 1996), which is also responsible for the recruitment of monocytes during chronic inflammation and therefore plays an important role in the transition from acute to chronic inflammation (Gabay 2006). IL1 receptor antagonist (IL1Ra) is secreted by various types of cells including immune cells, epithelial cells, and adipocytes, and is a natural inhibitor of the proinflammatory effect of IL1 $\beta$ (Perrier et al. 2006). Nevertheless, induction of IL1Ra production by other cytokines and acute phase proteins indicates a possible role of this cytokine in chronic inflammation (Arend et al. 1998). Moreover, recent studies implicated IL1Ra in the promotion of type 2 diabetes (Saltevo et al. 2008, Herder et al. 2009), another well-known risk factor for endometrial cancer development (Friberg et al. 2007).

We conducted a case-control study nested within the European Prospective Investigation into Cancer and nutrition (EPIC) to examine the relationships between endometrial cancer risk and blood concentrations of CRP, IL6, and IL1Ra. We also investigated to which extent these inflammatory markers can influence the association of obesity, endogenous sex steroids, and C-peptide (a marker of pancreatic insulin secretion) with endometrial cancer risk. With respect to endometrial cancer, this is the first prospective study of its kind to date, comprising 305 incident cases and 574 matched controls.

## Materials and methods

## Study population

The EPIC cohort is a large, multicenter prospective study, designed to investigate the associations between nutritional, lifestyle, metabolic, and genetic risk factors and cancer incidence. It was initiated in 1992 in ten European countries (Denmark, France, Germany, Greece, Italy, The Netherlands, Norway, Spain, Sweden, and the United Kingdom) and involved about 370000 women and 150000 men. About 246000 women and 140000 men also provided a blood sample. In the present study, Norway was not included because blood samples have been collected only recently, and very few cases of endometrial cancer have been diagnosed after blood donation.

Study population and baseline data collection have been previously described in details (Riboli et al. 2002). In brief, questionnaires included data about dietary, lifestyle and health factors, reproductive history, use of oral contraceptives (OCs) and hormone replacement therapy (HRT), history of any disorders or surgical operations, tobacco smoking and alcohol consumption, occupational history, physical activity, and education level.

## Collection and storage of blood samples

In France, The Netherlands, the United Kingdom, Germany, Spain, Italy, and Greece, blood samples were collected according to a standardized protocol (Riboli et al. 2002). From each subject, about 30 ml of blood was drawn, and serum, plasma, erythrocytes, and buffy coat were aliquoted in plastic straws of 0.5 ml each, which were heat-sealed and stored under liquid nitrogen ( $-196^{\circ} \mathrm{C}$ ) in a centralized biobank. In Denmark, blood fractions were aliquoted into 1 ml tubes, and stored in the vapor phase of liquid nitrogen containers ( $-150{ }^{\circ} \mathrm{C}$ ). In the Swedish center of Umea, blood samples were divided into ten aliquots of 1.5 ml each (six plasma, two buffy coat, and two erythrocytes), which were rapidly frozen at $-80^{\circ} \mathrm{C}$ in standard freezers.

## Determination of menopausal status at blood donation

Women were considered as premenopausal when they reported menstruating regularly over the past 12 months or, if data were incomplete, when they were $<42$ years of age at recruitment (99.5\% of EPIC women who had complete data were premenopausal below age 42). Women were considered as postmenopausal when they reported not having had any menses over the past 12 months, or when they reported bilateral ovariectomy. Women who had incomplete data were considered postmenopausal if they were older than 55. Women who had less than nine menstrual cycles during the year preceding recruitment or women aged between 42 and 55 years and with missing or incomplete questionnaire data or who reported the use of HRT were classified as having peri-menopausal or unknown menopausal status and excluded from the analyses.

## Follow-up for cancer incidence and vital status

Incident cancer cases were identified through several methods, including record linkage with regional cancer registries (Denmark, Sweden, Italy, The Netherlands, Spain, and the United Kingdom), health insurance
records, cancer and pathology registries, and active follow-up of study subjects (France, Germany, and Greece). Data on vital status were obtained from mortality registries at the regional or national level, in combination with data collected by active follow-up (Greece). For each EPIC center, closure dates of the study period were defined as the latest dates of complete follow-up for both cancer incidence and vital status (dates varied between centers, from June 1999 to December 2003).

## Selection of case and control subjects

Case subjects were selected among women who developed epithelial endometrial cancer after their recruitment into the cohort and before the end of the study period. Cases were coded according to the 10th revision of the International Statistical Classification of Injuries, Disease, and Deaths. Women whose menopausal status was classified as peri-menopausal/ unknown or who used any HRT at the time of blood donation, or any exogenous hormones for contraception or medical purposes, or those with a previous diagnosis of cancer (except non-melanoma skin cancer) were excluded from the study. Women whose endometrial cancers were not primary cancers, or who had a diagnosis of non-epithelial tumors, or who reported hysterectomy were also excluded.

A total of 305 incident cases of epithelial endometrial cancer were identified ( 80 cases among women who were premenopausal at blood donation and 225 cases among women who were postmenopausal at blood donation). The 305 incident cases included 52 cases in Denmark, 49 in Italy, 36 in Spain, 31 in the United Kingdom, 26 in The Netherlands, 8 in Greece, 13 in France, 18 in Germany, and 72 in Sweden.

For each case subject with endometrial cancer, two control subjects were chosen at random among appropriate risk sets consisting of all cohort members alive and free of cancer (except non-melanoma skin cancer) at the time of diagnosis of the index case. An incidence density sampling protocol for control selection was used, such that controls could include subjects who became a case later in time, while each control subject could also be sampled more than once. Matching characteristics for cases and controls included study recruitment center, menopausal status (premenopausal and postmenopausal), age ( $\pm 6$ months) at enrollment, time of the day of blood collection ( $\pm 1 \mathrm{~h}$ ), fasting status ( $<3 \mathrm{~h} ; 3-6 \mathrm{~h}$, and $>6 \mathrm{~h}$ ), and, for premenopausal women, phase of menstrual cycle ('follicular', 'peri-ovulatory', and 'luteal'). As described previously (Kaaks et al. 2005),
the menstrual cycle phase was determined using either backward or forward dating, depending on the study center. In centers where both dating methods were used, backward dating was preferred because the length of the luteal phase of the cycle is generally more constant than the length of the follicular phase. With forward dating, the first day of a woman's last period was set to 0 days, and all subsequent days were counted sequentially up to a maximum of 39 days. With backward dating, the first day of a woman's next period after blood donation was set to 28 days, and the date on which a blood sample had been provided was counted backwards from this date to a minimum acceptable value of -11 days. The follicular phase corresponds to days 0 to 11 with forward dating and days -11 to 11 with backward dating; the periovulatory phase corresponds to days 12 to 16 with both backward and forward dating; the luteal phase corresponds to days 17 to 39 with forward dating and days 17 to 28 with backward dating. For the present analysis, 574 controls were selected.

All participants had given their consent for their participation into the EPIC study. The Internal Review Board of IARC and local institutional review boards in participating centers have approved the study.

## Laboratory assays

Serum (or EDTA plasma for Swedish subjects) from cases and matched controls was analyzed within the same analytical batch. CRP and IL6 were measured by using enzyme-linked immunoassays (R\&D Systems Europe, Lille, France) at the International Agency for Research on Cancer (Section of Nutrition and Metabolism), while IL1Ra was measured by beadbased immunoassay (Linco, Millipore, Billerica, MA, USA) at the German Cancer Research Center (Division of Cancer Epidemiology). The different assays were chosen on the basis of a comparative study that was published previously (Dossus et al. 2009b).

For quality control, samples from two sera were analyzed in duplicate within each analytical batch. Mean intra-batch and inter-batch coefficients of variation, calculated on the concentrations from the quality control samples, were 6.8 and $9.3 \%$ for CRP, 6.3 and $8.2 \%$ for IL6, and 15 and $27.7 \%$ for IL1Ra. Samples from cases and controls, which were matched together, were analyzed within the same analytical batch. Twenty-two (2\%) CRP and 428 (50\%) IL1Ra values were below the limit of quantification. These values were set to the limit of quantification ( $16 \mathrm{pg} / \mathrm{ml}$ for IL1Ra and $78 \mathrm{ng} / \mathrm{ml}$ for CRP). The remaining missing values (due to problems with analyses or
limited amount of sample available) represented $<3 \%$ of the samples. Less than $2 \%$ of the samples had values above the highest calibration point for each marker. Owing to these random missing values, CRP was available for 290 complete matched sets of cases and controls (i.e. matched sets with CRP levels available for the case and at least one of the two controls), IL6 for 291, and IL1Ra for 301 complete sets.

Endogenous C-peptide and sex-steroid hormone measurements were performed on a substantial proportion of the study population ( 225 complete sets) at the International Agency for Research on Cancer, using commercially available immunoassays, and have been described in details elsewhere (Cust et al. 2007, Allen et al. 2008). In brief, serum testosterone, DHEAS, androstenedione, estradiol $\left(\mathrm{E}_{2}\right)$, and estrone $\left(\mathrm{E}_{1}\right)$ were measured by RIAs (Immunotech, Marseille, France and DSL, Webster, TX, USA), while C-peptide and sex hormone-binding globulin (SHBG) were measured by IRMAs (Immunotech and CIS-Bio, Gif-sur-Yvette, France). Serum $E_{2}$ levels were measured only in postmenopausal women because of its large intraindividual variation during the menstrual cycle among premenopausal women.

## Statistical analyses

In all analyses, measurements of inflammatory markers were transformed using the natural logarithm to normalize their distributions.

Pearson's partial correlation coefficients between inflammatory markers and with lifestyle factors were calculated among controls, adjusting for age at blood donation (continuous) and laboratory batch. Univariate associations of baseline characteristics, including BMI (continuous), age at first full-term pregnancy (continuous), number of full-term pregnancies (continuous), age at menarche (continuous), age at menopause (continuous), past OC (yes/no) or HRT use (yes/no), smoking status (never, former, and current smoker), alcohol consumption (continuous), physical activity (inactive, moderately inactive, moderately active, and active), and history of diabetes (yes/no), were evaluated using conditional logistic regression to retain the matching.

Odds ratios (ORs) for endometrial cancer in relation to BMI or circulating inflammatory marker levels were calculated by conditional logistic regression models. IL6 and CRP were categorized into quartiles, and the cut-off points were based on the distributions of the controls. Because about half of the IL1Ra values were below the limit of quantification of the assay, quartiles could not be used for this marker. Therefore,
undetectable IL1Ra values were grouped in one category used as the reference category, and values above the limit of quantification were categorized into tertiles (cut-off points based on the distributions of the controls with detectable values). BMI was categorized using the WHO definition of normal weight ( $<25 \mathrm{~kg} / \mathrm{m}^{2}$ ), overweight ( $25-29 \mathrm{~kg} / \mathrm{m}^{2}$ ), and obese ( $\geq 30 \mathrm{~kg} / \mathrm{m}^{2}$ ). Analyses on association between BMI and endometrial cancer risk included only complete matching sets (one case and at least one control) with data on C-peptide, $\mathrm{E}_{1}$, and inflammatory markers (258 cases and 452 controls). Likelihood ratio tests were used to assess linear trends in ORs with assigned quantitative scores $1,2,3$, and 4 for the categories. The effects of potential confounders (additionally to the matching criteria, controlled for by design) were examined by including additional terms into the logistic regression models. Potential confounders included BMI, waist circumference, age at menarche, age at menopause (among postmenopausal women only), previous use of OCs, previous use of HRT (among postmenopausal women only), smoking, alcohol consumption, physical activity, and diabetes. Only BMI and waist circumference affected point estimates by more than $10 \%$. Further adjustments for C-peptide, SHBG, testosterone, DHEAS, androstenedione, $\mathrm{E}_{1}$, and postmenopausal $\mathrm{E}_{2}$ were also performed.
For analysis of statistical heterogeneity between study countries or between subgroups of menopausal status, age at diagnosis, BMI categories, or lag time between blood donation and diagnosis, ORs were estimated for continuous measurements of inflammatory markers transformed on the $\log _{2}$ scale. In this scale, a unit increase corresponds to a doubling of concentration. Formal tests of heterogeneity between the ORs in different EPIC subgroups were based on $\chi^{2}$-statistics, calculated as the deviations of logistic $\beta$-coefficients observed in each of the subgroups, relative to the overall $\beta$-coefficient.
All statistical tests and corresponding $P$ values were two-sided, and $P$ values $<0.05$ were considered statistically significant. All analyses were performed using the SAS software package (Version 9, SAS Institute, Cary, NC, USA).

## Results

Baseline characteristics of cases and controls are presented in Table 1. On average, women diagnosed with endometrial cancer had a higher BMI than controls ( 27.5 vs $26.0, P<0.0001$ ), were more often nulliparous ( 16.9 vs $7.7 \%, P<0.0001$ ), were older at menopause ( 50.9 vs 49.8 years, $P=0.002$ ), were less

Table 1 Baseline characteristics of endometrial cancer cases and matched controls, means (s.D.) or percentages

| Variable | Cases ( $N=305$ ) | Controls ( $N=574$ ) | $P^{\text {a }}$ |
| :---: | :---: | :---: | :---: |
| Menopausal status |  |  | Matched |
| Premenopausal | 26.2\% | 25.1\% |  |
| Postmenopausal | 73.8\% | 74.9\% |  |
| Age at blood donation (years) | 56.9 (7.3) | 57.1 (7.4) | Matched |
| Age at diagnosis (years) | 60.4 (7.2) | - | - |
| Lag time between blood collection and diagnosis (years) | 3.5 (2.6) | - | - |
| Body mass index ( $\mathrm{kg} / \mathrm{m}^{2}$ ) | 27.5 (5.5) | 26.0 (4.3) | <0.0001 |
| Obese | 26.2\% | 16.4\% | 0.0002 |
| Waist circumference (cm) | 87.0 (13.1) | 83.4 (11.3) | <0.0001 |
| Nulliparous | 16.9\% | 7.7\% | <0.0001 |
| Number of full-term pregnancies ${ }^{\text {b }}$ | 2.3 (1.0) | 2.4 (1.1) | 0.09 |
| Age at first full-term pregnancy (years) ${ }^{\text {b }}$ | 24.7 (4.2) | 24.9 (4.3) | 0.96 |
| Age at menarche (years) | 13.0 (1.5) | 13.3 (1.6) | 0.10 |
| Age at menopause (years) ${ }^{\text {c }}$ | 50.9 (4.0) | 49.8 (3.9) | 0.002 |
| Previous OC use | 32.8\% | 42.0\% | 0.002 |
| Previous HRT use ${ }^{\text {c }}$ | 29.5\% | 16.7\% | 0.0001 |
| Self-reported diabetes | 4.6\% | 3.7\% | 0.50 |
| Physical activity |  |  | 0.64 |
| Inactive | 26.4\% | 27.1\% |  |
| Moderately inactive | 34.6\% | 37.8\% |  |
| Moderately active | 23.4\% | 19.5\% |  |
| Active | 15.6\% | 15.6\% |  |
| Alcohol consumption (g/day) | 7.1 (10.6) | 7.3 (10.3) | 0.86 |
| Smoking status |  |  | 0.18 |
| Never smokers | 66.5\% | 60.7\% |  |
| Former smokers | 19.9\% | 22.3\% |  |
| Current smokers | 13.5\% | 17.0\% |  |

${ }^{\text {a }} P$ values were calculated by using conditional logistic regression.
${ }^{\mathrm{b}}$ Among parous women only.
${ }^{\text {c} A m o n g ~ p o s t m e n o p a u s a l ~ w o m e n ~ o n l y . ~}$
often past OC users ( 32.8 vs $42.0 \%, P=0.002$ ), and more often past HRT users ( 29.5 vs $16.7 \%, P=0.0001$ ) as reported previously (Dossus et al. 2009a).

Differences in cytokine levels according to baseline characteristics were evaluated among controls after adjusting for age at blood donation, laboratory batch, and BMI. Age at menopause, past HRT and OC use, self-reported diabetes, physical activity, and alcohol consumption were not significantly associated with any of the inflammatory markers (data not shown). However, higher CRP and IL6 mean levels were observed among current smokers (CRP: $1528 \mathrm{ng} / \mathrm{ml}$ and IL6: $1.61 \mathrm{pg} / \mathrm{ml}$ ) compared to past smokers (CRP: $1066 \mathrm{ng} / \mathrm{ml}$ and IL6: $1.38 \mathrm{pg} / \mathrm{ml}$ ) and never smokers (CRP: $1066 \mathrm{ng} / \mathrm{ml}$ and IL6: $1.22 \mathrm{pg} / \mathrm{ml}$ ). IL6 mean levels were also higher in premenopausal women $(1.36 \mathrm{pg} / \mathrm{ml})$ than in postmenopausal women $(1.13 \mathrm{pg} / \mathrm{ml})$. After adjustment for age at blood donation and laboratory batch, there was a relatively strong correlation between IL6 and CRP levels ( $r=0.51,95 \%$ confidence interval (CI): 0.44-0.57; Table 2). Lower degrees of correlation were observed between IL1Ra and IL6 ( $r=0.17,95 \%$ CI: $0.09-0.25$ )
or between IL1Ra and CRP ( $r=0.23,95 \%$ CI: $0.15-0.30$ ). Pearson partial correlations with age at blood donation varied between 0.10 for IL1Ra ( $95 \%$ CI: 0.01-0.18), 0.18 for CRP ( $95 \%$ CI: $0.10-0.26$ ), and 0.28 for IL6 ( $95 \%$ CI: $0.20-0.35$ ). All markers were correlated with BMI ( $r=0.37$, $95 \%$ CI: $0.29-0.44$ for CRP; $r=0.32$, $95 \%$ CI: $0.24-0.39$ for IL6; and $r=0.18,95 \%$ CI: $0.10-0.26$ for IL1Ra) and waist circumference ( $r=0.43$, $95 \%$ CI: $0.35-0.50$ for CRP; $r=0.35,95 \%$ CI: $0.27-0.43$ for IL6; and $r=0.24,95 \%$ CI: $0.15-0.32$ for IL1Ra). Scatter plots of BMI by inflammatory marker levels for cases and controls are presented in Supplementary Figure 1, see section on supplementary data given at the end of this article. None of the markers correlated with time before centrifugation or storage time of the samples ( $r<0.05$, data not shown).

Conditional logistic regression analyses showed an increased risk of endometrial cancer with the three inflammatory markers studied, with an OR of 1.58 ( $95 \%$ CI: 1.03-2.41) for the highest versus the lowest quartile of CRP, 1.66 ( $95 \%$ CI: $1.08-2.54$ ) for the highest versus the lowest quartile of IL6, and 1.82

Table 2 Pearson's partial correlation coefficients ( $95 \%$ confidence interval (CI)) for associations between inflammatory factors, body mass index (BMI), waist circumference, and age at blood donation among controls

|  | CRP | IL6 | IL1Ra |
| :--- | :---: | :---: | :---: |
| L6 $^{\text {a }}$ | $0.51(0.44-0.57)$ |  |  |
| LL1Ra $^{\mathrm{a}}$ | $0.23(0.15-0.30)$ | $0.17(0.09-0.25)$ |  |
| BMI $^{\mathrm{a}}$ | $0.37(0.29-0.44)$ | $0.32(0.24-0.39)$ | $0.18(0.10-0.26)$ |
| Waist circumference $^{\mathrm{a}}$ | $0.43(0.35-0.50)$ | $0.35(0.27-0.43)$ | $0.24(0.15-0.32)$ |
| Age at blood donation $^{\mathrm{b}}$ | $0.18(0.10-0.26)$ | $0.28(0.20-0.35)$ | $0.10(0.01-0.18)$ |

${ }^{\text {a }}$ Analyses on log-transformed data, adjusted for age at blood donation and laboratory batch.
${ }^{\mathrm{b}}$ Analyses on log-transformed data, adjusted for laboratory batch.
( $95 \%$ CI: 1.22-2.73) for the highest versus the lowest category of IL1Ra (Table 3). After adjustment for BMI, these estimates were lower (from - $25 \%$ for IL6 to $-33 \%$ for CRP) and were no longer statistically significant. A similar reduction in risk estimates was also observed after adjustment for waist circumference ( $-19 \%$ for IL1Ra and $-30 \%$ for IL6 and CRP) (data not shown). Further adjustment for C-peptide resulted in an additional risk reduction of $6-12 \%$, whereas further adjustment for $\mathrm{E}_{1}$ only reduced risk estimate by $0-5 \%$ (Table 4). In postmenopausal women, we observed a slightly stronger effect of the adjustment for $\mathrm{E}_{1}$ on the association between endometrial cancer and CRP (OR in the highest quartile changed from $1.36,95 \%$ CI: $0.72-2.59$ to $1.21,95 \%$ CI: $0.63-2.34$ ) or IL6 (OR in the highest quartile changed from 1.23, $95 \%$ CI: $0.67-2.25$ to $1.14,95 \% \mathrm{CI}: 0.62-2.11$ ). Very similar estimates were observed after adjustment for postmenopausal $\mathrm{E}_{2}$ ( OR in the highest quartile: $1.23,95 \%$ CI: $0.63-2.40$ for CRP and 1.12, $95 \%$

CI: 0.60-2.10 for IL6). Adjustments for testosterone, androstenedione, DHEAS, or SHBG did not change risk estimates by more than $5 \%$ (data not shown). Exclusion of subjects with CRP $\geq 10 \mathrm{mg} / \mathrm{l}$ (a level indicative of acute inflammation) did not affect risk estimates (data not shown).
Since half of the subjects had values below the limit of quantification for IL1Ra, we performed sensitivity analyses excluding these subjects. After adjustment for BMI, ORs were 1.23 ( $95 \%$ CI: $0.61-2.49$ ) for women in the second tertile and 1.92 ( $95 \%$ CI: 1.02-3.64) for women in the third tertile compared to women in the first tertile of measured IL1Ra ( $P_{\text {trend }}=0.04$ ).
No heterogeneity was observed in the associations of CRP and IL1Ra with endometrial cancer risk by country, menopausal status (premenopausal versus postmenopausal), age at diagnosis ( $<55$ vs $\geq 55$ years), BMI categories ( $<25,25-29$, and $\geq 30 \mathrm{~kg} / \mathrm{m}^{2}$ ), or lag time between blood donation and diagnosis ( $<2$ vs $\geq 2$ years and $<5$ vs $\geq 5$ years) (data not shown).

Table 3 Risk (odds ratio, OR ( $95 \%$ confidence interval, CI )) of endometrial cancer by categories of inflammatory markers

|  | Categories |  |  |  | $P_{\text {trend }}{ }^{\text {a }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 |  |
| CRP |  |  |  |  |  |
| Quartile cut-offs ( $\mathrm{ng} / \mathrm{ml})^{\text {b }}$ | <568 | 568-1154 | 1155-2232 | >2232 |  |
| Cases/controls | 61/136 | 65/135 | 77/136 | 87/135 |  |
| Crude (matched) | 1.00 | 1.11 (0.73-1.70) | 1.34 (0.89-2.03) | 1.58 (1.03-2.41) | 0.02 |
| Adjusted for BMI | 1.00 | 0.99 (0.65-1.53) | 1.14 (0.74-1.75) | 1.05 (0.65-1.69) | 0.70 |
| IL6 |  |  |  |  |  |
| Quartile cut-offs (pg/ml) ${ }^{\text {b }}$ | <0.85 | 0.85-1.21 | 1.22-1.89 | >1.89 |  |
| Cases/controls | 64/135 | 57/132 | 75/133 | 95/134 |  |
| Crude (matched) | 1.00 | 0.94 (0.61-1.45) | 1.27 (0.83-1.96) | 1.66 (1.08-2.54) | 0.008 |
| Adjusted for BMI | 1.00 | 0.84 (0.54-1.31) | 1.05 (0.67-1.65) | 1.24 (0.78-1.98) | 0.22 |
| IL1Ra |  |  |  |  |  |
| Category cut-offs (pg/ml) ${ }^{\text {b }}$ | $\leq 16.0^{\text {c }}$ | 16.1-59.6 | 59.7-140.7 | > 140.7 |  |
| Cases/controls | 142/286 | 33/91 | 51/91 | 75/93 |  |
| Crude (matched) | 1.00 | 0.80 (0.50-1.27) | 1.28 (0.83-1.97) | 1.82 (1.22-2.73) | 0.004 |
| Adjusted for BMI | 1.00 | 0.71 (0.44-1.14) | 1.09 (0.70-1.70) | 1.44 (0.94-2.21) | 0.10 |

[^1]Table 4 Risk (odds ratio, OR ( $95 \%$ confidence interval, CI )) of endometrial cancer by categories of inflammatory markers, in women with C -peptide and estrone measurements

|  | Categories |  |  |  | $P_{\text {trend }}{ }^{\text {a }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 |  |
| CRP |  |  |  |  |  |
| Quartile cut-offs ( $\mathrm{ng} / \mathrm{ml})^{\text {b }}$ | <564 | 564-1171 | 1172-2329 | >2329 |  |
| Cases/controls | 43/99 | 54/97 | 63/99 | 65/98 |  |
| Adjusted for BMI | 1.00 | 1.16 (0.71-1.92) | 1.25 (0.76-2.06) | 1.06 (0.61-1.85) | 0.75 |
| Adjusted for BMI and C-peptide | 1.00 | 1.04 (0.62-1.73) | 1.07 (0.64-1.79) | 0.93 (0.53-1.64) | 0.83 |
| Adjusted for BMI and estrone | 1.00 | 1.17 (0.71-1.93) | 1.27 (0.77-2.10) | 1.01 (0.58-1.77) | 0.86 |
| IL6 |  |  |  |  |  |
| Quartile cut-offs (pg/ml) ${ }^{\text {b }}$ | <0.85 | 0.85-1.19 | 1.20-1.80 | >1.80 |  |
| Cases/controls | 46/99 | 45/97 | 60/99 | 74/98 |  |
| Adjusted for BMI | 1.00 | 0.91 (0.55-1.51) | 1.16 (0.69-1.97) | 1.33 (0.78-2.30) | 0.20 |
| Adjusted for BMI and C-peptide | 1.00 | 0.87 (0.52-1.45) | 1.06 (0.62-1.82) | 1.19 (0.68-2.07) | 0.40 |
| Adjusted for BMI and estrone | 1.00 | 0.93 (0.56-1.56) | 1.17 (0.69-2.00) | 1.30 (0.75-2.25) | 0.25 |
| IL1Ra |  |  |  |  |  |
| Category cut-offs (pg/ml) ${ }^{\text {b }}$ | $\leq 16.0^{\text {c }}$ | 16.1-49.5 | 49.6-120.2 | >120.2 |  |
| Cases/controls | 106/211 | 24/59 | 41/61 | 54/62 |  |
| Adjusted for BMI | 1.00 | 0.80 (0.45-1.43) | 1.24 (0.74-2.06) | 1.56 (0.94-2.58) | 0.08 |
| Adjusted for BMI and C-peptide | 1.00 | 0.79 (0.44-1.40) | 1.10 (0.65-1.86) | 1.47 (0.88-2.45) | 0.16 |
| Adjusted for BMI and estrone | 1.00 | 0.83 (0.47-1.48) | 1.24 (0.74-2.06) | 1.56 (0.94-2.60) | 0.08 |

${ }^{a} P$ value for trend with assigned quantitative scores $1,2,3$, and 4 for the categories.
${ }^{\mathrm{b}}$ Cut-off points were based on the distribution of controls (of subjects with detectable values).
${ }^{\text {c }}$ Values below the limit of quantification.

For IL6, a stronger risk was observed for a doubling of concentrations among premenopausal women (OR: $1.50,95 \% \mathrm{CI}: 1.00-2.24$ ) compared to postmenopausal women (OR: 1.09, $95 \%$ CI: 0.88-1.34; Fig. 1). A similar heterogeneity was observed between cases diagnosed after or before 2 years following blood donation (OR: $1.43,95 \%$ CI: 1.00-2.04 for cases diagnosed $<2$ years after blood donation and OR: $1.07,95 \%$ CI: $0.86-1.32$ for cases diagnosed more than 2 years after blood donation). However, the heterogeneity tests were not statistically significant ( $P_{\text {heterogeneity }}=0.16$ and 0.17 ). Exclusion of cases diagnosed in the first 2 years of follow-up did not modify the heterogeneity observed by menopausal status. Similarly, the heterogeneity observed between early and late diagnosed cases persisted when the analysis was restricted to postmenopausal women.

An OR of 2.02 ( $95 \%$ CI: $1.26-3.23, P_{\text {trend }}=0.005$ ) was observed for obese women (BMI $\geq 30 \mathrm{~kg} / \mathrm{m}^{2}$ ) compared to normal weight women (BMI $<25 \mathrm{~kg} / \mathrm{m}^{2}$; Table 5). Adjustments for inflammatory markers resulted in reduction of risk estimates by $11 \%$ for CRP, $15 \%$ for IL6, and $18 \%$ for IL1Ra. After adjustment for C-peptide, obese women had an OR of 1.43 ( $95 \% \mathrm{CI}: 0.86-2.40, P_{\text {trend }}=0.21$ ) compared to normal weight women. Adjustment for $\mathrm{E}_{1}$ reduced the OR for obese versus normal weight women to 1.75 ( $95 \%$ CI: 1.07-2.85, $P_{\text {trend }}=0.03$ ). Further adjustment
for CRP, IL6, or IL1Ra resulted in a reduction in risk estimates for obese women ( -3 to $-18 \%$ ). In the models adjusted for C-peptide, $\mathrm{E}_{1}$, and inflammatory markers, the ORs for obese versus normal weight women were close to unity (OR: 1.21, $95 \% \mathrm{CI}$ : $0.69-2.14$ with further adjustment for CRP; OR: 1.16, 95\% CI: 0.67-2.01 with further adjustment for IL6; OR: $1.10,95 \% \mathrm{CI}: 0.64-1.90$ with further adjustment for IL1Ra). A similar reduction in the estimates was observed for the risk of endometrial cancer associated with waist circumference after adjustment for inflammatory markers, $\mathrm{E}_{1}$, and C-peptide (data not shown).

## Discussion

In this prospective study, we observed that circulating levels of CRP, IL6, and IL1Ra were significantly associated with endometrial cancer risk, although the association was largely dependent on the levels of adiposity. The association between adiposity measurements and endometrial cancer was also substantially attenuated after adjustment for inflammatory markers, even when the effect of C-peptide or $\mathrm{E}_{1}$ was already taken into account.

One mechanism for the inflammation-mediated association between obesity and endometrial cancer could be the modulation of aromatase activity by cytokines within the adipose tissue (Purohit \& Reed 2002).


Figure 1 Risk (OR ( $95 \%$ CI)) of endometrial cancer for IL6 on a continuous $\log _{2}$ scale adjusted for BMI, and stratified by menopausal status and lag time to cancer diagnosis.

After menopause, when the estrogen production from the ovary has ceased, most of the circulating estrogens derive from the peripheral aromatase conversion of androgens in the adipose tissue. It has been shown that IL6 can stimulate aromatase activity in the adipose tissue (Zhao et al. 1995) and therefore enhance the estrogen production and bioavailability. This hypothesis is also supported by the fact that adjustment for inflammatory markers, in addition to $\mathrm{E}_{1}$, further attenuated the risk among obese women.

Furthermore, inflammatory markers have also been shown to play an important role in the development of insulin resistance, hyperglycemia, and type 2 diabetes (Greenberg \& McDaniel 2002), which are known risk factors for endometrial cancer (Kaaks et al. 2002). IL6 is one of the major cytokines released by the adipose tissue (Orban et al. 1999, Fain et al. 2004). Its concentration correlates with measures of obesity and insulin resistance (Yudkin et al. 1999, Vozarova et al. 2001), and with the development of type 2 diabetes (Pradhan et al. 2001, Spranger et al. 2003, Hu et al. 2004). Moreover, IL6 enhances the hepatic production of the suppressor of cytokine signaling-3 (a protein that is known to inhibit the insulin signaling pathway) (Lebrun \& Van 2008) and decreases glucose uptake in the adipose tissue (Bastard et al. 2002, Feve \& Bastard 2009). IL6 is also the main regulatory factor of hepatic production of CRP, which also decreases insulin signaling and glucose uptake (Devaraj et al. 2009). IL1Ra is another cytokine that is released in major amounts by the adipose tissue (Juge-Aubry et al. 2003)
and is elevated more than sevenfold in obesity (Meier et al. 2002). More recently, this cytokine has also been implicated in the development of type 2 diabetes (Saltevo et al. 2008, Herder et al. 2009). Its role in obesity-related inflammation and in the development of associated diseases is not yet fully understood. However, it has been hypothesized that IL1Ra might have a role in the development of resistance to leptin in obese patients, and this in turn may further sustain obesity (Luheshi et al. 1999, Meier et al. 2002). IL1Ra could also contribute to an increase in adiposity by inhibiting IL1 $\beta$, which inhibits lipogenesis on the one hand, while stimulates lipolysis, glucose transport, and adipocyte maturation in the adipose tissue on the other hand (Juge-Aubry et al. 2003).

It is also possible that cytokines have a direct effect on endometrial carcinogenesis. Within the uterus, IL6 and IL1 are present throughout the menstrual cycle, with a peak of expression during the proliferative phase (Kelly et al. 2001, von Wolff et al. 2002). Cytokines have been shown to enhance the growth and metastasis of various types of tumors (Aggarwal et al. 2006), and alterations in the levels of cytokine production have been reported in several types of cancer, including gynecologic cancers. For example, increased IL6 concentrations have been reported in patients with endometrial carcinoma (Chopra et al. 1997, Punnonen et al. 1998, Bellone et al. 2005, Slater et al. 2006), and laboratory experiments have demonstrated that nuclear factor- $\kappa$ B (NF- $\kappa \mathrm{B}$ ), a cellular transcription factor that activates genes for immune and inflammatory response, is aberrantly expressed in a majority of endometrial cancer tumors (Vaskivuo et al. 2002, Pallares et al. 2004). NF-кB activation also leads to COX-2 expression, which induces elevated levels of prostaglandin $\mathrm{E}_{2}$, a protein that has been shown to promote the transformation of normal endometrium into neoplastic tissue (Jabbour et al. 2006).

The major strengths of our study are the high degree of standardization across countries for blood collection protocols and questionnaire data and the fact that blood samples were collected prospectively prior to endometrial cancer diagnosis. This design reduces the possibility of 'reverse causation bias', which may occur when the presence of a tumor modifies circulating inflammatory marker levels. Indeed, the stronger association observed for IL6 among cases diagnosed in the 2 years following blood collection suggests the presence of a preclinical tumor that can release IL6 (Scambia et al. 1994, Bellone et al. 2005, Slater et al. 2006) and may have influenced circulating levels. However, this was not observed for the other two markers. Another possible explanation for

Table 5 Risk (odds ratio, OR ( $95 \%$ confidence interval, CI)) of endometrial cancer by categories of body mass index (BMI), with further adjustment for inflammatory markers, C-peptide, and estrone

|  | BMI WHO categories $\left(\mathrm{kg} / \mathrm{m}^{2}\right)$ |  |  |  |
| :--- | :--- | :--- | :--- | :--- |
|  | $<\mathbf{2 5}$ | $\mathbf{2 5 - 2 9}$ | $\mathbf{3 0 +}$ | $\boldsymbol{P}_{\text {trend }}{ }^{\text {a }}$ |
| Cases/controls | $81 / 174$ | $82 / 151$ | $61 / 68$ |  |
| Crude (matched) | 1.00 | $1.23(0.82-1.84)$ | $2.02(1.26-3.23)$ | 0.005 |
| Adjusted for estrone | 1.00 | $1.14(0.75-1.72)$ | $1.75(1.07-2.85)$ | 0.03 |
| Adjusted for C-peptide | 1.00 | $1.02(0.67-1.56)$ | $1.43(0.86-2.40)$ | 0.21 |
| Adjusted for C-peptide and estrone | 1.00 | $0.94(0.61-1.46)$ | $1.24(0.73-2.12)$ | 0.49 |
| Adjusted for CRP | 1.00 | $1.15(0.76-1.76)$ | $1.79(1.06-3.02)$ | 0.04 |
| Adjusted for estrone and CRP | 1.00 | $1.08(0.70-1.66)$ | $1.57(0.92-2.69)$ | 0.12 |
| Adjusted for C-peptide and CRP | 1.00 | $1.00(0.65-1.55)$ | $1.39(0.80-2.41)$ | 0.29 |
| Adjusted for C-peptide, estrone, and CRP | 1.00 | $0.93(0.60-1.46)$ | $1.21(0.69-2.14)$ | 0.57 |
| Adjusted for IL6 | 1.00 | $1.14(0.76-1.73)$ | $1.72(1.04-2.85)$ | 0.04 |
| Adjusted for estrone and IL6 | 1.00 | $1.06(0.70-1.62)$ | $1.51(0.90-2.53)$ | 0.15 |
| Adjusted for C-peptide and IL6 | 1.00 | $0.99(0.64-1.52)$ | $1.33(0.78-2.27)$ | 0.35 |
| Adjusted for C-peptide, estrone, and IL6 | 1.00 | $0.92(0.59-1.43)$ | $1.16(0.67-2.01)$ | 0.68 |
| Adjusted for IL1Ra | 1.00 | $1.15(0.76-1.73)$ | $1.65(1.01-2.71)$ | 0.05 |
| Adjusted for estrone and IL1Ra | $1.07(0.70-1.62)$ | $1.45(0.87-2.41)$ | 0.18 |  |
| Adjusted for C-peptide and IL1Ra | $0.99(0.64-1.51)$ | $1.26(0.74-2.14)$ | 0.44 |  |
| Adjusted for C-peptide, estrone, and IL1Ra | 1.00 | $0.91(0.59-1.42)$ | $1.10(0.64-1.90)$ | 0.80 |

${ }^{\text {a }} P$ value for trend with assigned quantitative scores 1,2 , and 3 for the categories.
the stronger association with IL6 in the first years of follow-up could be that IL6 levels at recruitment may not be predictive of inflammatory status after several years of follow-up during which major determinants of IL6 levels such as smoking status, physical activity, or medical treatments might have changed.

One limitation of our study is that only one cytokine measurement, at a single point in time, was available to reflect long-term exposure to chronic inflammation. This could result in random misclassification, in view of intra-individual variations over time, and therefore may have led to underestimation of risk estimates. However, a recent study measuring cytokine levels in repeated blood samples taken over a 2 -year period showed moderate to good reliability with correlations in the range of $0.57-0.92$ for the inflammatory markers included in our study ( Gu et al. 2009). The latter results suggest that, while random misclassification certainly occurs, its attenuating effects on risk estimates are moderate. Another limitation is the fact that about half of the values for IL1Ra were below the limit of quantification of the assay. Results for this marker should therefore be considered with caution until reproduced in other studies.

In conclusion, elevated prediagnostic levels of the inflammatory markers CRP, IL6, and IL1Ra were associated with an increased risk of endometrial cancer. Our results provide some degree of support to the hypothesis that chronic inflammation, through its
association with insulin resistance and estrogen production, but also independently of these two pathways, might mediate the obesity-related increase in risk of endometrial cancer, although after adjustment for BMI the associations were weaker and no longer statistically significant. To our knowledge, this is the first study of its kind, and further prospective studies are needed to confirm these observations. Further experimental studies should also investigate whether inflammation has a direct and independent effect on endometrial carcinogenesis.

## Supplementary data

This is linked to the online version of the paper at http://dx. doi.org/10.1677/ERC-10-0053.

## Declaration of interest

The authors declare that there is no conflict of interest that could be perceived as prejudicing the impartiality of the research reported.

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[^1]:    ${ }^{\text {a }} P$ value for trend with assigned quantitative scores $1,2,3$, and 4 for the categories.
    ${ }^{\text {b }}$ Cut-off points were based on the distribution of controls (of subjects with detectable values).
    ${ }^{c}$ Values below the limit of quantification.

